

The skew-normal distribution and related multivariate families

Adelchi Azzalini

Università di Padova

NordStat 2004: Jyväskylä, 6–10 June 2004



Overview

- Background & motivation
- Skew-normal distribution (scalar/multivariate case, extensions)
- More general families
 - skew-elliptical families, and some special cases
 - a general formulation
 - ‘semi-parametric’ families
- Connection with other areas
 - selective sampling
 - stochastic frontier models
 - forms of robust likelihood
 - applications in mathematical geology, finance, etc.
- Some open problems



Some general remarks



Motivation

- Context: parametric families of pdf's for continuous variates
 - Want to build flexible parametric classes
 - extend normal family, in scalar and multivariate case
 - construct families of “arbitrary” flexibility
 - bridge between parametric and non-parametric approach
 - normal class is central to multivariate data analysis, ... but
- “who has ever seen a multivariate normal sample?”

[Barnett (1979)]



An old history

- (bivariate) Pearson family
- marginal transformation approach, including:
 - Johnson family
 - g - and h -transformations
- Edgeworth expansions
- mixtures and compound
- some approaches are specific of d -dimensional case:
 - copulae
 - specification via conditional distributions
- others...

[Reviews: Pretorius (1930), Johnson & Kotz (1970–72), Hutchinson & Lay (1990), Joe (1997), Patil et al. (1976), Kotz et al. (2000→), ...]



Desiderata

- high flexibility in shape with wide range for indices of skewness and kurtosis
- mathematically tractable:
 - simple form of density, moments, cdf, etc.
 - closed under marginalization and conditioning
 - retain nice properties of original family
 - parameters should be directly linked to some aspects of pdf
- association with ‘physical’ mechanism(s)
 - to support statistical modeling
 - for random number generation
- allow easy and reliable inference



Expanding the Normal family: an approach



A simple yet useful result

Lemma (1D case)

Suppose f_0 and G' are pdf's on $\mathbb{R}$ symmetric about 0, then

$$f(z) = 2 f_0(z) G\{w(z)\}$$

is a density for any odd function $w(\cdot)$.

Proof. If $Y \sim f_0$ and $X \sim G'$ are independent,

$$\begin{aligned} \frac{1}{2} &= \mathbb{P}\{X - w(Y) \leq 0\} = \mathbb{E}_Y \{\mathbb{P}\{X - w(Y) \leq 0 | Y = z\}\} \\ &= \int_{\mathbb{R}} G\{w(z)\} f_0(z) \, dz. \end{aligned}$$

[Azzalini (SJS, 1985) for case $w(z) = \alpha z$]



The skew-normal distribution ($d = 1$)

Use lemma to ‘skew’ the $N(0,1)$ density:

$$\phi(z; \alpha) = 2 \underbrace{\phi(z)}_{f_0(z)} \underbrace{\Phi(\alpha z)}_{G\{w(z)\}}$$

for $z, \alpha \in \mathbb{R}$. Write

$$Z \sim \text{SN}(\alpha)$$

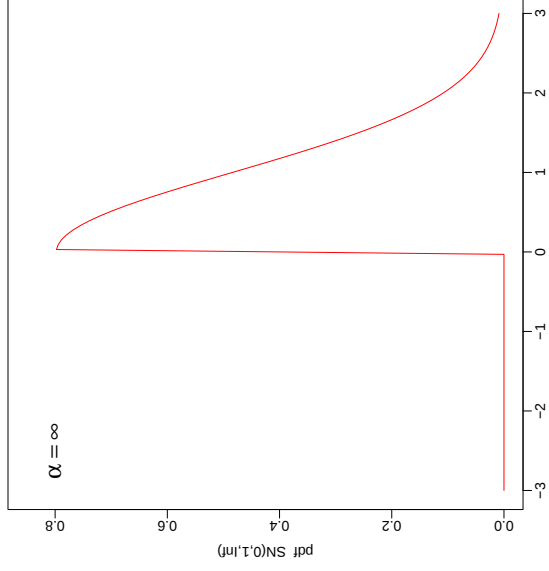
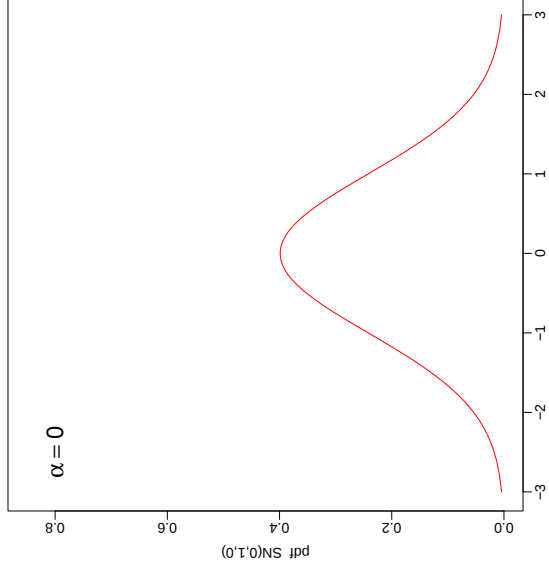
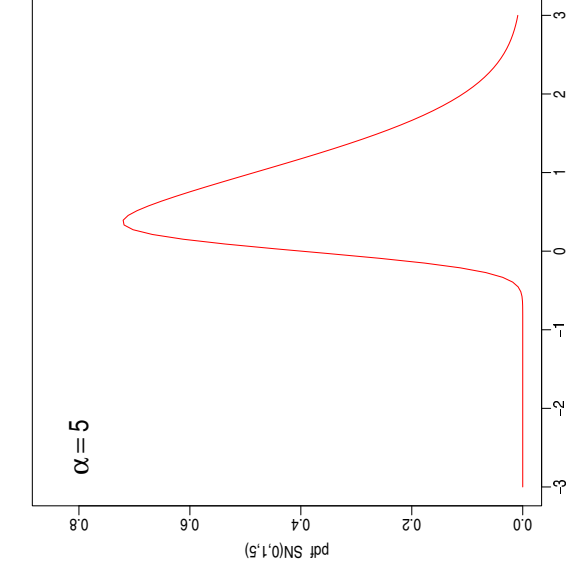
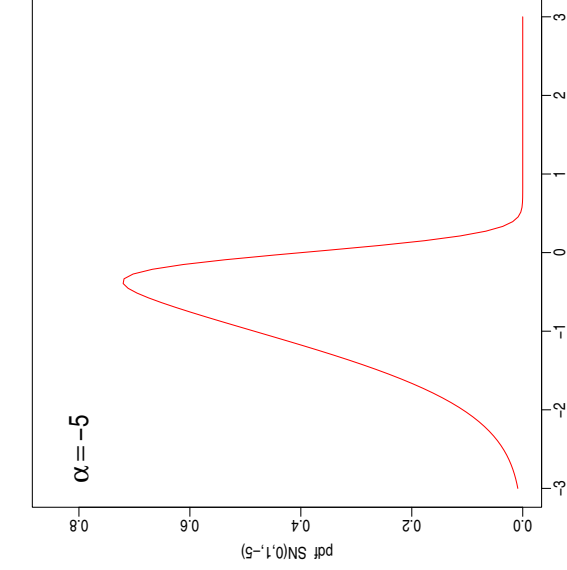
Some basic facts:

- if $\alpha = 0$, obtain $N(0,1)$
- if $Z \sim \text{SN}(\alpha)$, then $-Z \sim \text{SN}(-\alpha)$
- $1 - \Phi(-z; \alpha) = \Phi(z; -\alpha)$ for its cdf



‘Extended’ version: $\phi(z) \Phi(\alpha_0 + \alpha z) / \Phi(\alpha_0 / \sqrt{1 + \alpha^2})$

Some SN pdf's ($d = 1$)



Forms of genesis

- Conditioning:
if $(U_0, U_1) \sim N_2$ with $N(0,1)$ marginals and correlation ρ ,
then

$$U_0|U_1 > 0 \sim \text{SN} \left(\rho/\sqrt{1-\rho^2} \right)$$

For $U_0|U_1 > \tau \sim ESN$

- Sum of Normal+HN:
if V_0, V_1 iid $N(0, 1)$, then

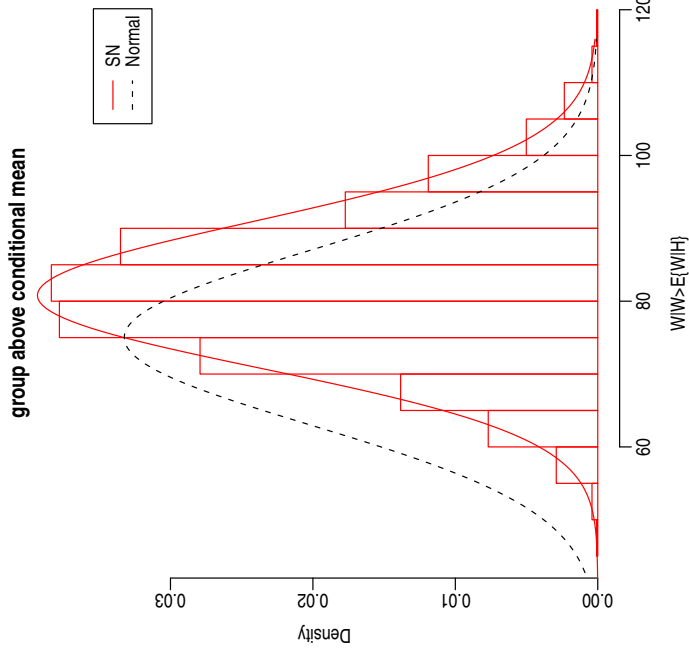
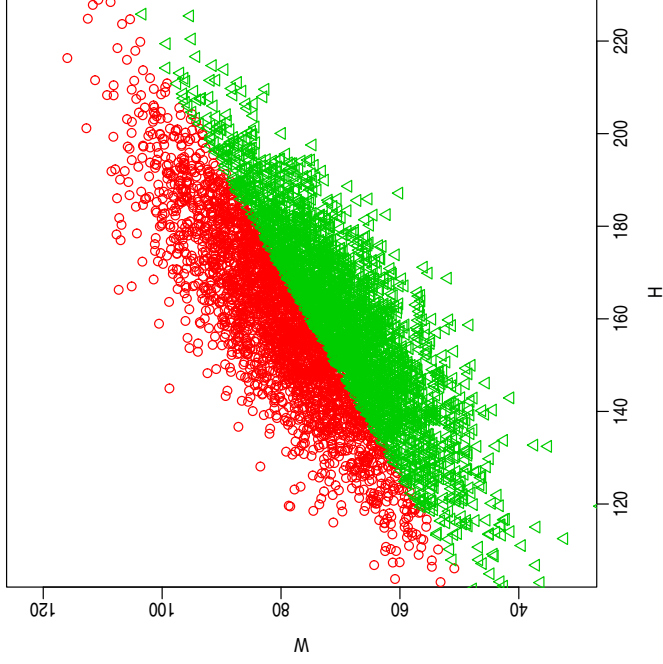
$$\delta |V_0| + \sqrt{1-\delta^2} V_1 \sim \text{SN} \left(\delta/\sqrt{1-\delta^2} \right)$$

- Order statistics:

$$\max(U_0, U_1) \sim \text{SN} \left(\sqrt{(1-\rho)/(1+\rho)} \right)$$



A “constructive” example



$$\begin{pmatrix} H \\ W \end{pmatrix} \sim N_2 \left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} V_{11} & V_{12} \\ V_{21} & V_{22} \end{pmatrix} \right)$$

separate out group with W

‘above average’ (given H):

$$W|U > 0, \quad U = W - \mathbb{E}\{W|H\}$$

$$W \sim N(\mu_2, V_{22})$$

$$(W|U > 0) \sim \text{SN}(\mu_2, V_{22}, \delta/\sqrt{1 - \delta^2})$$

$$\delta = \sqrt{1 - \rho^2}$$



Some nice properties

- $\mathbb{P}\{Z < z\} = \Phi(z) - 2T(z, \alpha)$, T is Owen's function
- $Z^2 \sim \chi_1^2$, independent of α
- hence even moments are the same of $N(0, 1)$
- odd moments from N^* HN representation
- or use $\mathbb{E}\left\{e^{tZ}\right\} = 2e^{t^2/2}\Phi(\delta t)$, $\delta = \alpha/\sqrt{1 + \alpha^2}$
- $SN(\alpha)^*N(0, 1) \sim$ (scaled) $SN(\alpha/\sqrt{2 + \alpha^2})$
- etc...

[Azzalini (1985), Henze (1986), Chiogna (1998)]



Early (independent) appearances

- Conditioning:
Birnbaum (1950) obtains pdf of $U_1|U_0 > \tau$
and some properties of first two moments.
Motivation: selection process of personnel, in education, etc.
- Sum of Norm+HN:
 - in *Technom.* (1964), discussion with contributions by Weinstein, Lipow, Mantel, Wilkinson and Nelson
 - in Bayesian context, O'Hagan & Leonard (1976)
obtain ESN from a two-stage prior construction
- Order statistics:
Roberts (1966) obtains SN pdf, and its property

$$Z^2 \sim \chi_1^2$$



Other ways of skewing the normal pdf

- use previous lemma with $G \neq \Phi$, many options...

[Nadarajah & Kotz (2003)]



$$f(z) = \text{constant} \times \phi(z) \Phi(\alpha z)^n, \quad n \in \mathbb{N}$$

connection with order statistics of $N(0,1)$

[Balakrishnan (in discussion on Arnold & Beaver, 2002)]

- or take entirely different routes: two-pieces normal, etc.



Skewing non-normal distributions

$$f(z) = 2 f_0(z) G(\alpha z), \quad f_0(z) \neq \phi(z)$$

- f_0 exponential power dist'n (or GED or Normal^(r)) for heavy (or light) tails $\Rightarrow$ SEP

[Azzalini (1986), DiCiccio & Monti (2004)]

- skew-Laplace [Balakrishnan and Ambalagaspitya (1994)]
- skew-Cauchy [Arnold & Beaver (2000b), include multivariate version]
- skew-logistic [Wahed & Ali (2000)]
- the wrapped SN on the circle (related to above form but not quite the same)

[Pewsey (2000b, 2003)]

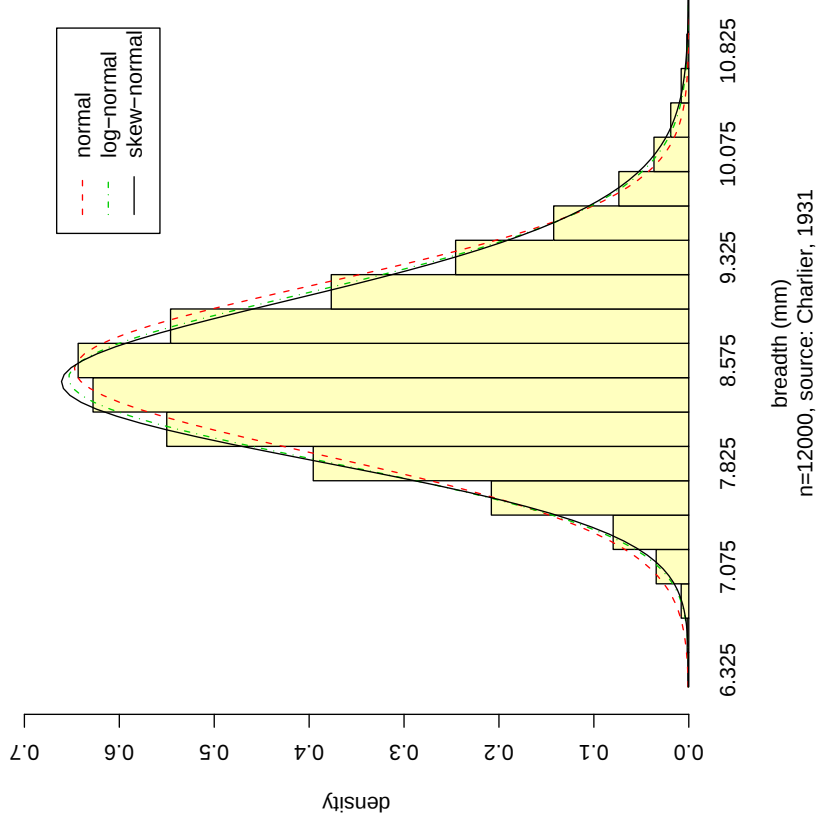


Statistical aspects



An old example: Johannsen's beans data

Breadth of Phaseolus Vulgaris
Johannsen's data



Distribution	X^2	d.f.	p-value
Normal	196.5	13	< 0.001
Log-normal	28.1	13	0.008
Edgeworth			
1st order	34.3	12	< 0.001
2nd order	14.9	11	0.19
Skew-normal	14.2	12	0.29

Note: data re-grouped from 20 into 16 bins
[Charlier (1931), Cramér (1946, p. 440–1)]

$$\hat{\gamma}_1 = 0.2809, \quad \hat{\gamma}_1 \sqrt{n/6} = 12.56$$



Likelihood function is ‘regular’, but...

- at $\alpha = 0$:
 - ◊ always a stationarity point of $\log L$
 - ◊ Fisher expected information matrix becomes singular

[connection with non-standard cases of Rotnitzky et al. (2000)]

Both can be avoided by suitable reparametrization

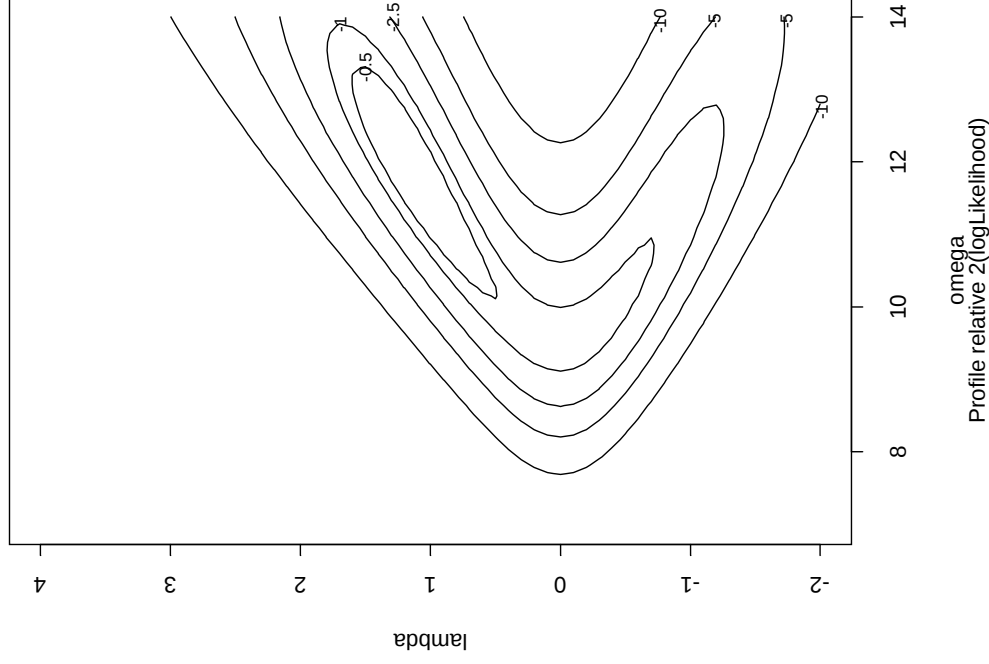
- MLE can occur at $\alpha = \pm\infty$ for ‘small n ’
some solutions proposed
- MLE can have more than one local maximum
- ‘small’ here is larger than elsewhere

[Azzalini (1985), Liseo (1990), Azzalini & Capitanio (1999), Pewsey (2000a), Monti (2003), Sartori (2003+)]

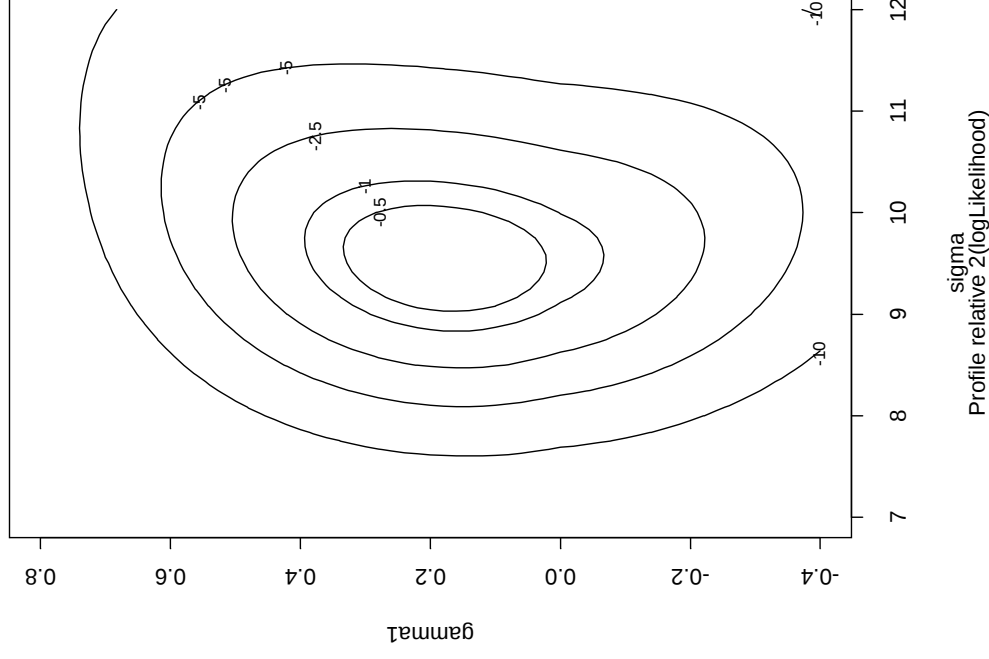


An example of profile log-likelihood

dataset: otis



dataset: otis



use of ‘direct’ vs ‘centred’ parameterization



Bayesian inference

- ESN arises from two-stage prior

$$Y \sim N(\theta, \tau^2), \quad \theta \sim N(\mu, \sigma^2), \quad \mu \sim N_+(\mu_0, \sigma_0^2)$$

[O'Hagan & Leonard (1976)]

- Bayesian approach to estimation of parameters can avoid $\hat{\alpha} = \infty$.

[Liseo (1990), see also Liseo & Loperfido (2002, to appear)]

- linear Bayesian estimation for model

$$Y \sim N(\theta, 1), \quad \theta \sim \text{SN}(\lambda)$$

[Mukhopadhyay & Vidakovic (1995)]



Multivariate SN distribution



Multivariate versions of 1D genesis

- Conditioning:

$$\begin{pmatrix} U_0 \\ U_1 \end{pmatrix} \sim N_{d+1} \left(0, \begin{pmatrix} 1 & \delta^\top \\ \delta & \bar{\Omega} \end{pmatrix} \right), \quad N(0, 1) \text{ marginals}$$

$$(U_1 | U_0 > 0) \sim \text{SN}_d(0, \bar{\Omega}, \alpha), \quad \alpha = \text{function}(\delta, \bar{\Omega})$$

- Sum of Norm+HN:

$$V_0 \sim N(0, 1), \quad V_1 \sim N_d(0, \Psi), \quad V_0 \perp\!\!\!\perp V_1$$

$$Z_j = \delta_j |V_0| + \sqrt{1 - \delta_j^2} V_{1j}, \quad j = 1, \dots, d$$

$$Z \sim \text{SN}_d(0, \bar{\Omega}, \alpha), \quad (\bar{\Omega}, \alpha) = \text{function}(\delta, \Psi)$$

The two mechanisms generate the same class of pdf's



[Azzalini & Dalla Valle (1996)]

Multivariate SN distribution

‘Standard’ SN: write $Z \sim \text{SN}_d(0, \bar{\Omega}, \alpha)$ if pdf at $z \in \mathbb{R}^d$ is

$$f(z) = 2 \phi_d(z; \bar{\Omega}) \Phi(\alpha^\top z)$$

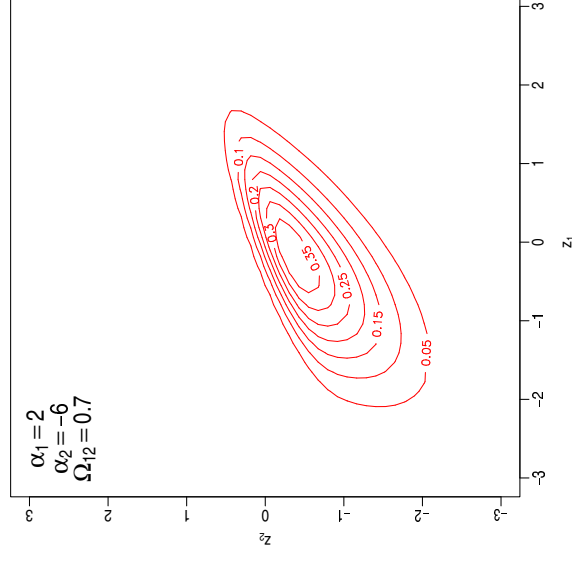
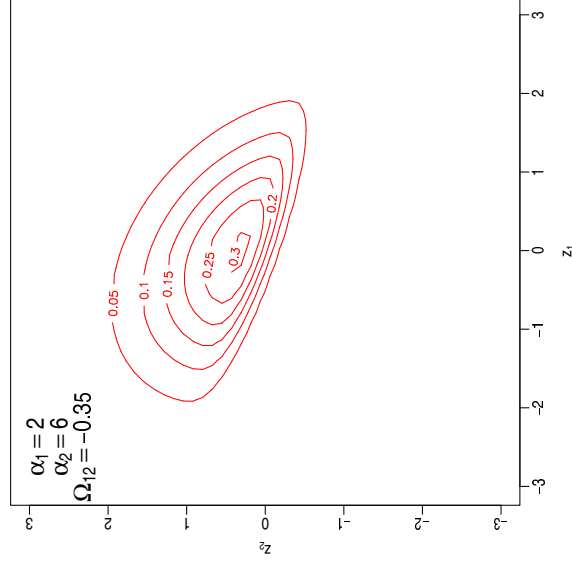
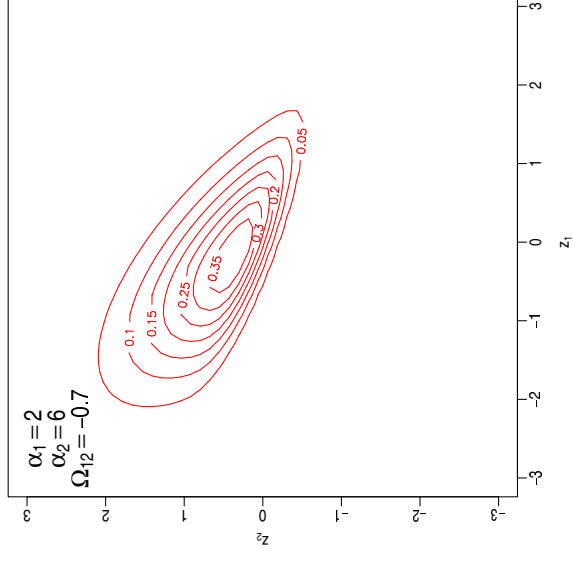
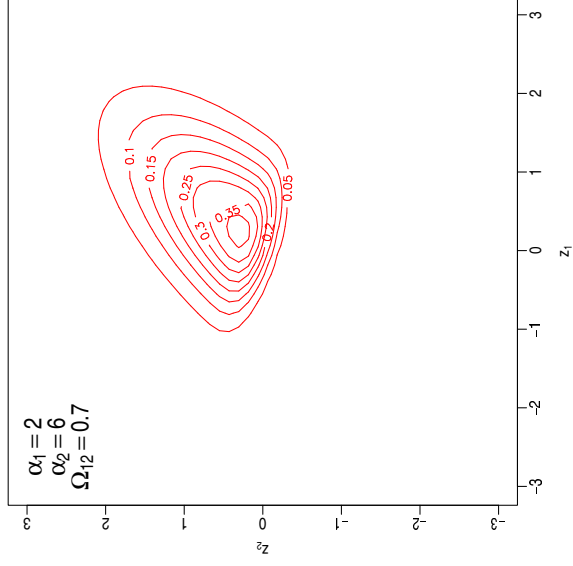
$\alpha \in \mathbb{R}^d$. If $\alpha = 0$, get $N_d(0, \bar{\Omega})$.

Add location/scale parameters:

$$Y = \xi + \omega Z \sim \text{SN}_d(\xi, \Omega, \alpha), \quad \Omega = \omega \bar{\Omega} \omega$$



Some SN₂ pdf's



Some formal properties – I

- Parameters ξ, Ω, α are variation-independent
- Formal properties depend on

$$\delta = (1 + \alpha^\top \bar{\Omega} \alpha)^{-1/2} \bar{\Omega} \alpha$$

e.g.

$$\mathbb{E}\{Z\} = \sqrt{2/\pi} \, \delta = \mu, \quad \text{var}\{Z\} = \bar{\Omega} - \mu \mu^\top$$

- moment generating function & cumulants:

$$\mathbb{E}\left\{e^{tZ}\right\} = 2 \exp\left(\frac{1}{2} t^\top \bar{\Omega} t\right) \Phi(\delta^\top t)$$

$$\implies k\text{-th order cumulant:} \quad \text{constant} \times \delta_r \delta_s \cdots \delta_v$$

$$\implies \gamma_{1,d} \in (0, 0.9905), \quad \gamma_{2,d} \in (0, 0.869)$$



Some formal properties – II

- Random numbers: any of
 - ◊ $Z = \text{sign}(U_0) U_1$
 - ◊ $Z = \delta |V_0| 1_d + (1 - \delta^2)^{1/2} V_1$
require a N_{d+1} variate
- cdf of SN_d computed via cdf of N_{d+1} , namely
$$\mathbb{P}\{Z < z\} = 2 \mathbb{P}\{-U_0 < 0 \cap U_1 < z\}$$
- the class is closed under affine transformations

$$a + AY \sim \text{SN}_p(a + A\xi, A\Omega A^\top, \alpha')$$

hence also closed under marginalisation



A characterisation property

A parallel of a well-known property of the Normal class.

Let X be a r. v. (in $\mathbb{R}^d$) with $\mathbb{E}\{X\} = \mu$ and $\text{var}\{X\} = \Sigma$.

Put $\Omega = \Sigma + \mu \mu^\top$. If

$$h^\top X \sim \text{SN}(\alpha_h)$$

for any h such that $h^\top \Omega h = 1$, then $X \sim \text{SN}_d$.

[Gupta & Huang (2002)]



Quadratic forms

- $(Y - \xi)^\top \Omega^{-1} (Y - \xi) \sim \chi_d^2$

leads to Healy-type diagnostic plot

- if $X \sim N_d(0, \Omega)$ and $Y \sim \text{SN}_d(0, \Omega, \alpha)$, then

$$X^\top A X \stackrel{d}{=} Y^\top A Y$$

- more general: for an ‘even’ function $T(\cdot)$, such that
 $T(x) = T(-x)$ for all $x \in \mathbb{R}^d$,

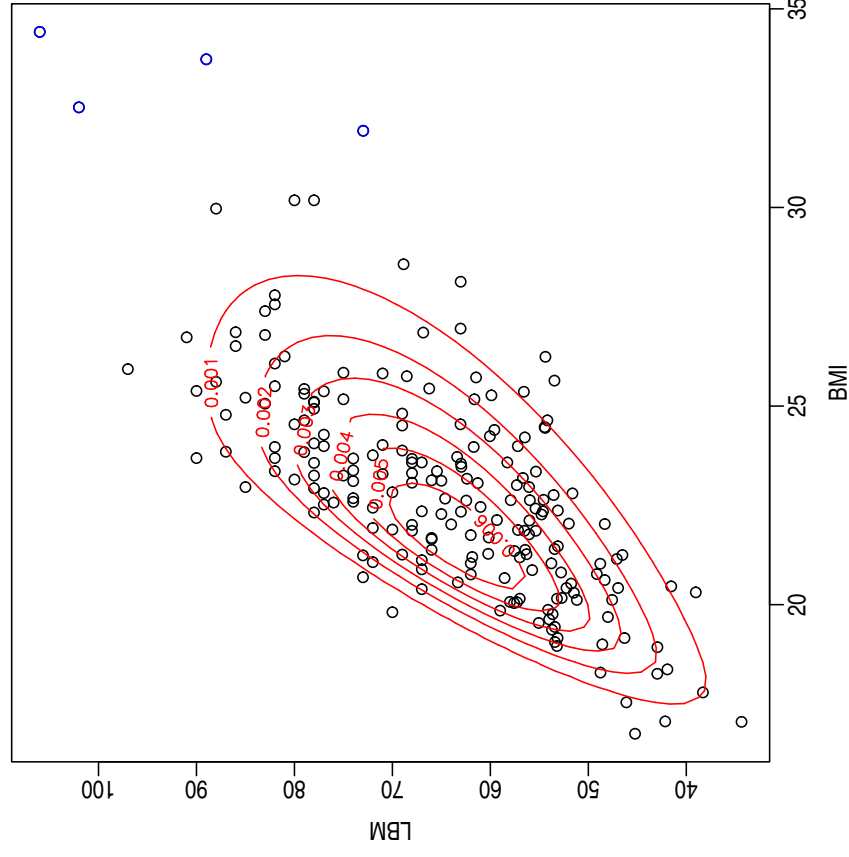
$$T(X) \stackrel{d}{=} T(Y)$$

[Azzalini & Capitanio (1999, 2003), Genton et al. (2001), Loperfido (2001), Genton & Loperfido (to appear)]

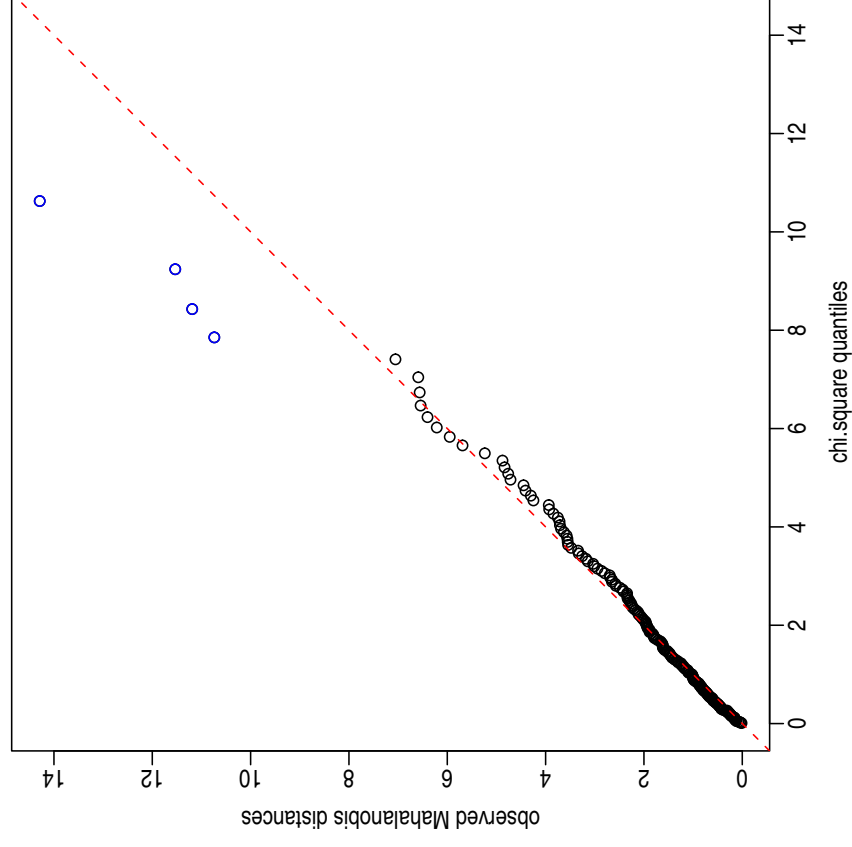


A numerical example: AIS data

Australian Institute of Sport data ($n = 202$)



scatterplot and SN fit using MLE



Healy's QQ plot for SN fit



ESN and graphical models

‘Extended’ SN: $Z = (U_1 | U_0 > \tau)$ with pdf

$$\text{const} \times \phi_d(z; \bar{\Omega}) \Phi(\alpha_0 + \alpha^\top z), \quad \alpha_0 = \tau \sqrt{1 + \alpha^\top \bar{\Omega} \alpha}$$

ESN class $\{Y = \xi + \omega Z\}$ is closed under conditioning

[Arnold & Beaver (2000a)]

Graphical models:

- can build a conditional independence graph for Y (with some restrictions)
- relationship between the conditional independence graphs of $\xi + \omega U \sim N_d(\xi, \Omega)$ and of $Y \sim \text{SN}_d(\xi, \Omega, \alpha, \tau)$
- log-likelihood allows a parameter based factorization



[Capitanio, Azzalini & Stanghellini (SJS, 2003)]

Stochastic processes

- spatial processes with SN components, application to rainfall data, Bayesian inference via MCMC [Kim & Mallick (2004)]
- model for non-Gaussian noise in signals processing [Gualtierotti (2004)]



Multiple constraints

Consider $Z = (U_1|U_0 > \tau)$ where $(U_0, U_1) \sim N_{h+d}$

Various cases:

- product of $\Phi(\cdot)$ terms [Azzalini (1985), Arnold & Beaver (2000a)]
- case $h = d$, involves a $\Phi_d(\cdot)$ term
[Sahu, Dey & Branco (2004)]
- Hierarchical SN [Liseo & Loperfido (2003)]
- Fundamental SN [Arellano-Valle & Genton (to appear)]
- General SN

$$\phi_d(y; \xi, \Omega) \Phi_h(D(y - \xi); \nu, \Delta) / \Phi_h(D\xi; \nu, \Delta + D\Omega D^\top)$$

[mentioned in Gupta, González-Farías & Domínguez-Molina (2004), with focus on case $\nu = 0$, $\Delta = I_h$]



Perturbation of symmetric multivariate distributions



Extend basic result to $\mathbb{R}^d$

(central) symmetry about 0 in $\mathbb{R}^d$: $f_0(y) = f_0(-y)$

N.B. includes elliptically contoured densities

Lemma. Suppose f_0 and G' are pdf's symmetric about 0 in $\mathbb{R}^d$ and $\mathbb{R}$, respectively, then

$$f(z) = 2 f_0(z) G\{w(z)\}, \quad z \in \mathbb{R}^d,$$

is a density for any real-valued odd function $w(\cdot)$, i. e.
 $w(-x) = -w(x) \in \mathbb{R}$ for $x \in \mathbb{R}^d$.

Proof. Essentially the same of case $d = 1$.

Can be extended to h -dimensional G , and other directions.



[Azzalini & Capitanio (2003), Genton & Loperfido (to appear),
see also Arellano-Valle, del Pino & San Martín (2002) for another formulation]

A simple stochastic representation

Suppose $X \sim G'$, $Y \sim f_0$, independent. Then

$$Z = \begin{cases} Y & \text{if } X < w(Y) \\ -Y & \text{if } X > w(Y) \end{cases}$$

has density

$$f(z) = 2 f_0(z) G\{w(z)\}.$$

Corollary. If $Y \sim f_0$ and $Z \sim f$, with $f(z) = 2 f_0(z) G\{w(z)\}$, then

$$T(Y) \stackrel{d}{=} T(Z)$$

for any even function $T(\cdot)$.



Can obtain high flexibility – an example

Consider a symmetric Beta pdf, scaled in $(-1, 1)^2$:

$$f_0(y) = \frac{(1 - y_1^2)^{a-1} (1 - y_2^2)^{b-1}}{4^{a+b-1} B(a, a) B(b, b)}, \quad y = (y_1, y_2) \in (-1, 1)^2,$$

perturbed by

$$G(x) = \frac{e^x}{1 + e^x}, \quad w(y) = \frac{\sin(p_1 y_1 + p_2 y_2)}{1 + \cos(q_1 y_1 + q_2 y_2)}$$

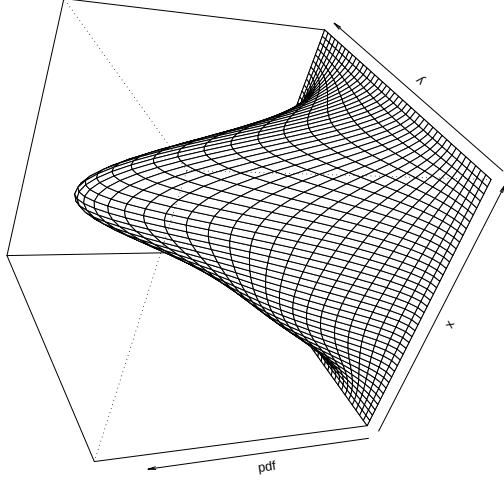
where p_1, p_2, q_1, q_2 are additional parameters.

There is more than skewing in the ‘skew-symmetric’ class

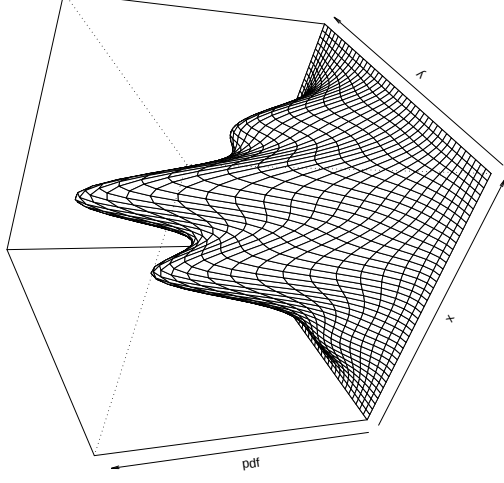


Perturbed symmetric Beta pdf's

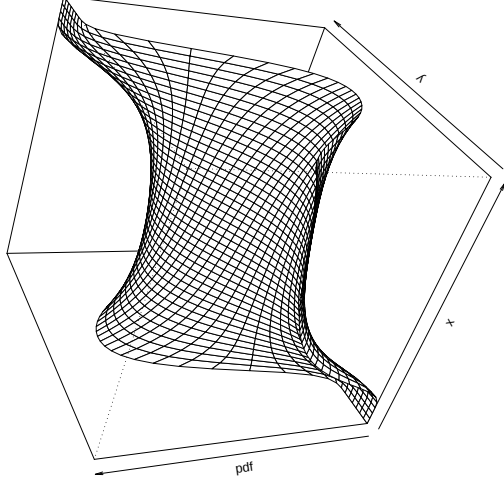
$$(a,b,p,q) = (2, 3, (3, 3), (0, 0))$$



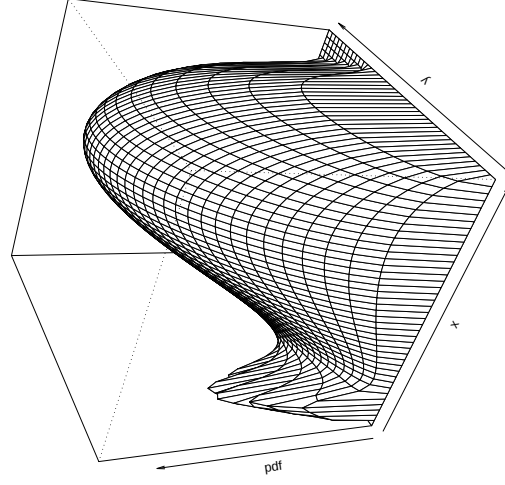
$$(a,b,p,q) = (2, 3, (8, 8), (0, 0))$$



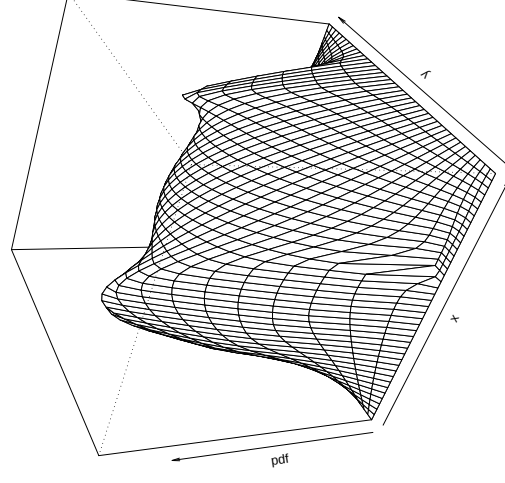
$$(a,b,p,q) = (3, 1, (-1, 3), (2, 1))$$



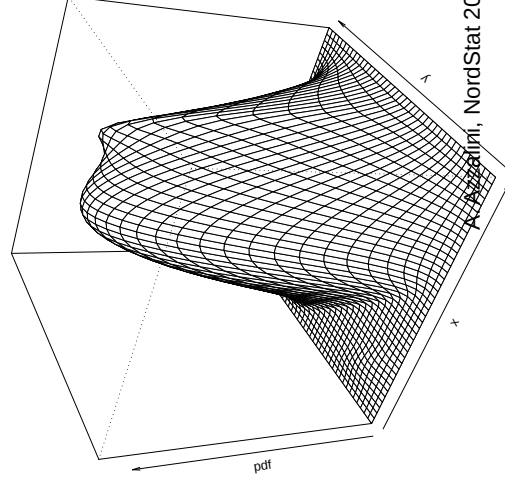
$$(a,b,p,q) = (3, 1.5, (3, 1), (2.5, 1))$$



$$(a,b,p,q) = (3, 2, (2, -3), (2, 4))$$



$$(a,b,p,q) = (3, 3, (1, 1), (3, 3))$$



Skew-elliptical densities

$$f_0(y) = \frac{c_d}{|\Omega|^{1/2}} \tilde{f} \left[(y - \xi)^\top \Omega^{-1} (y - \xi) \right]$$

- Route A: use previous lemma, with some $G(\cdot)$ and $w(\cdot)$
[Azzalini & Capitanio (1999)]
- Route B: apply a random scale factor to a SN variate
[Branco & Dey (2001)]

Good news: for appropriate choice of the ingredients,
‘essentially’ the same thing. [Azzalini & Capitanio (2003)]



Analogies of SN and S-elliptical families

- the two stochastic representations for SN variates still hold, replacing normal with elliptical variates
 - ◊ via conditioning: $(U_1|U_0 > 0)$,
 - ◊ via addition: $\delta|V_0|1_d + \sqrt{1 - \delta^2} V_1$
- for case $d = 2$, also representation via order statistics
- the distribution of quadratic forms is the same for skew-elliptical and elliptical variates with common symmetric component

[Azzalini & Capitanio (2003), Fang (2003)]



A noteworthy case: skew- t

If $Z \sim SN(0, \Omega, \alpha)$ and $V \sim \chi_\nu^2/\nu$, then

$$Y = \xi + V^{-1/2} Z$$

has pdf

$$f(y) = 2 t_d(y; \nu) T_1(w(y); \nu + d)$$

where

$$w(y) = \alpha^\top \omega(y - \xi) \left(\frac{d + \nu}{\|y - \xi\|_\Omega^2 + \nu} \right)^{1/2}.$$

Unlimited range of indices of skewness and kurtosis.
Suitable for robust inference.

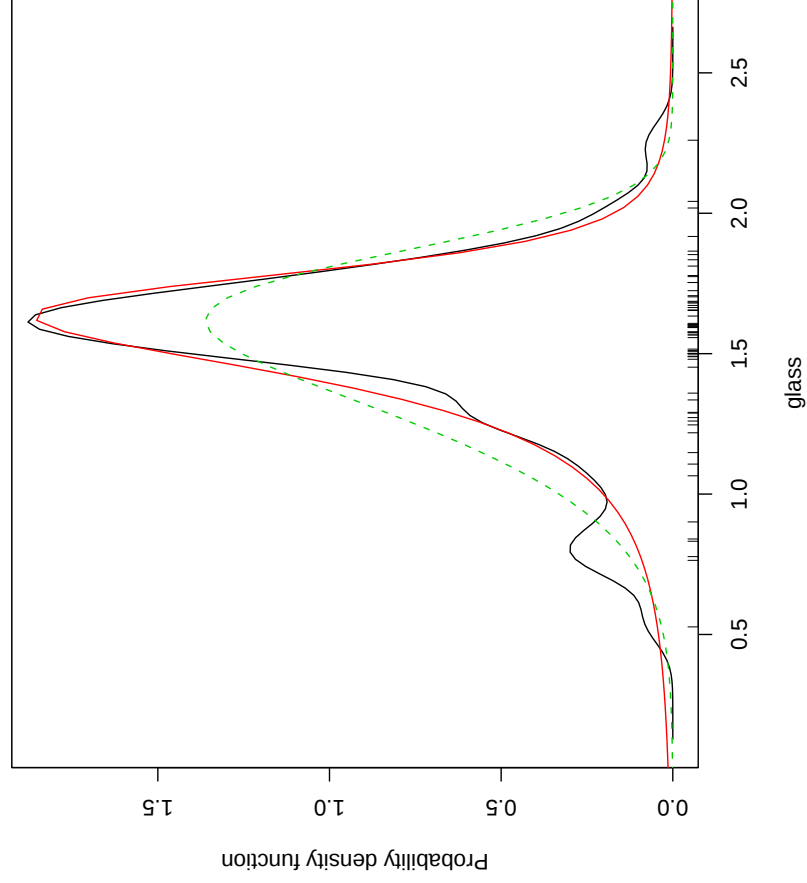
$$(Y - \xi)^\top \Omega^{-1} (Y - \xi) \sim d \times F(d, \nu) \quad \text{for QQ- and PP-plot}$$



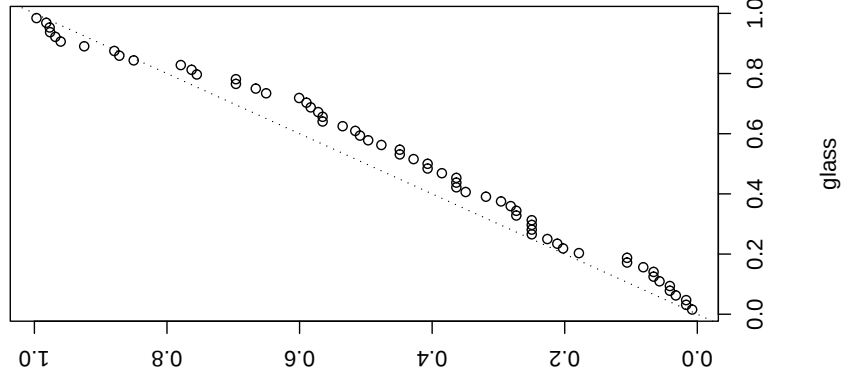
Example: fiber glass strength

data on breaking strength of fiber glass ($n = 63$)

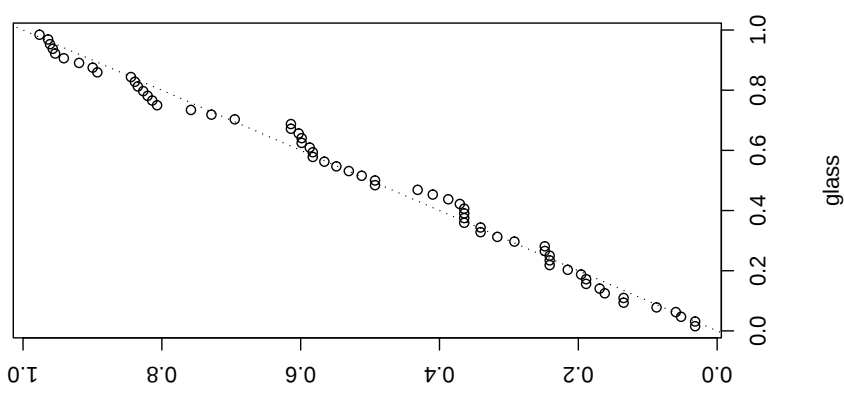
Glass-fiber data: nonparametric, SN and St fit



PP-plot for normal distribution



PP-plot for skew-t distribution



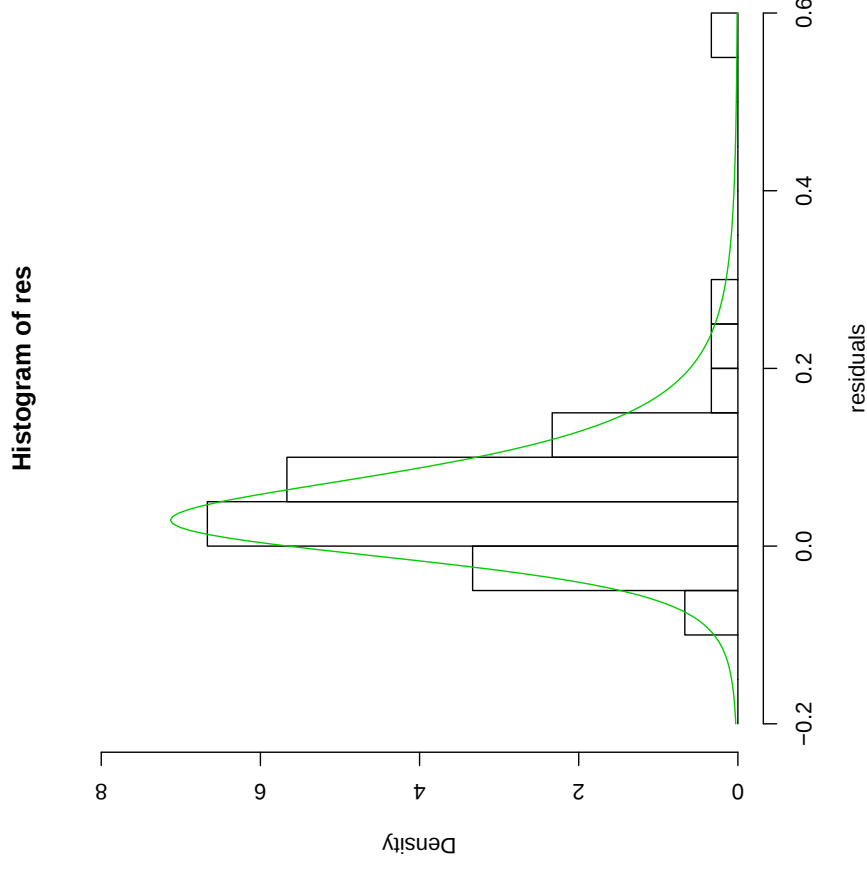
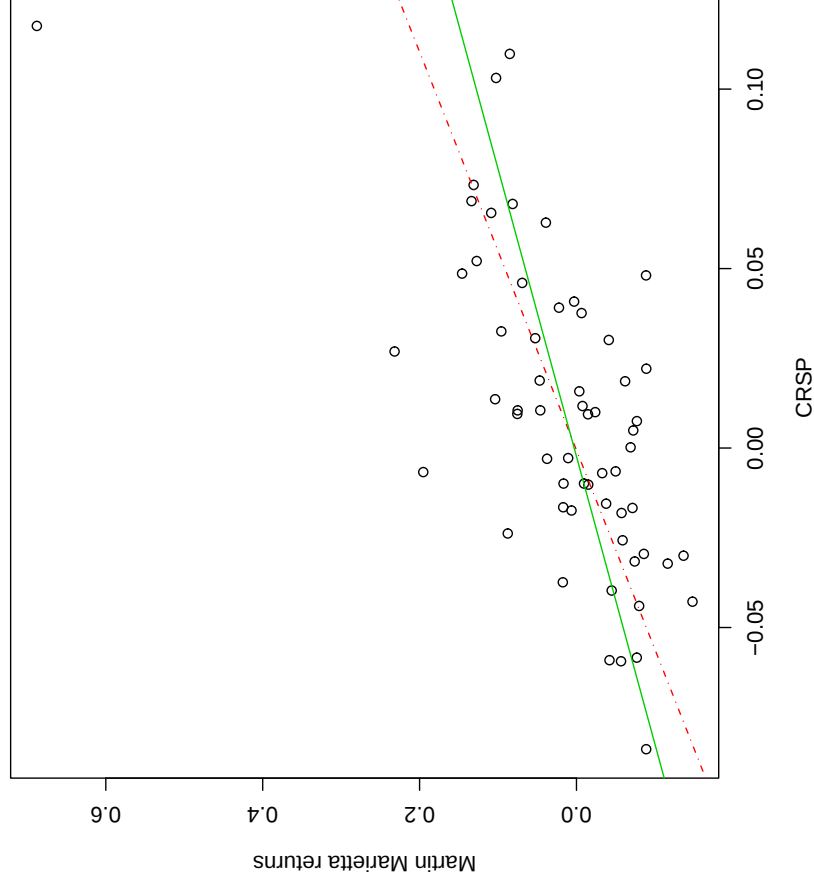
nonparametric estimate, SN fit, ST fit

PP-plot for Normal model and ST model



Example: MM returns

MM excess returns vs CRSP index ($n = 60$)



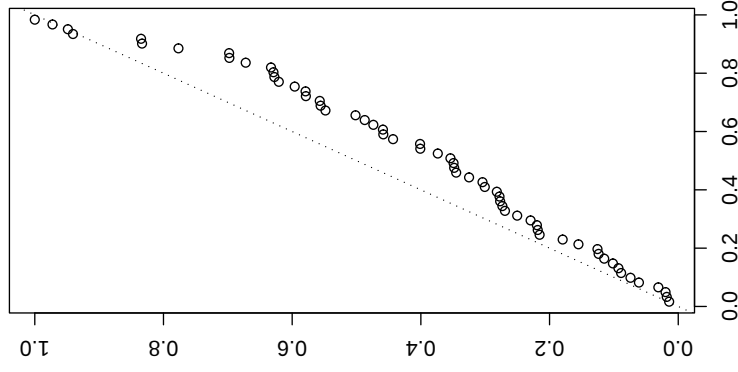
linear regression with Normal and ST fit

histogram of residuals and fitted ST dist'n

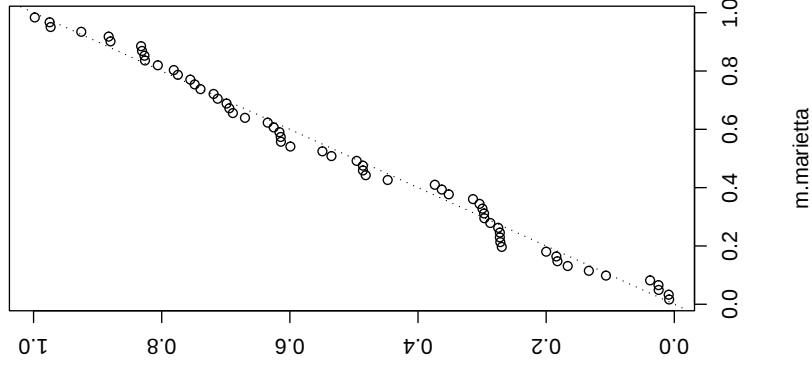


Example: MM returns

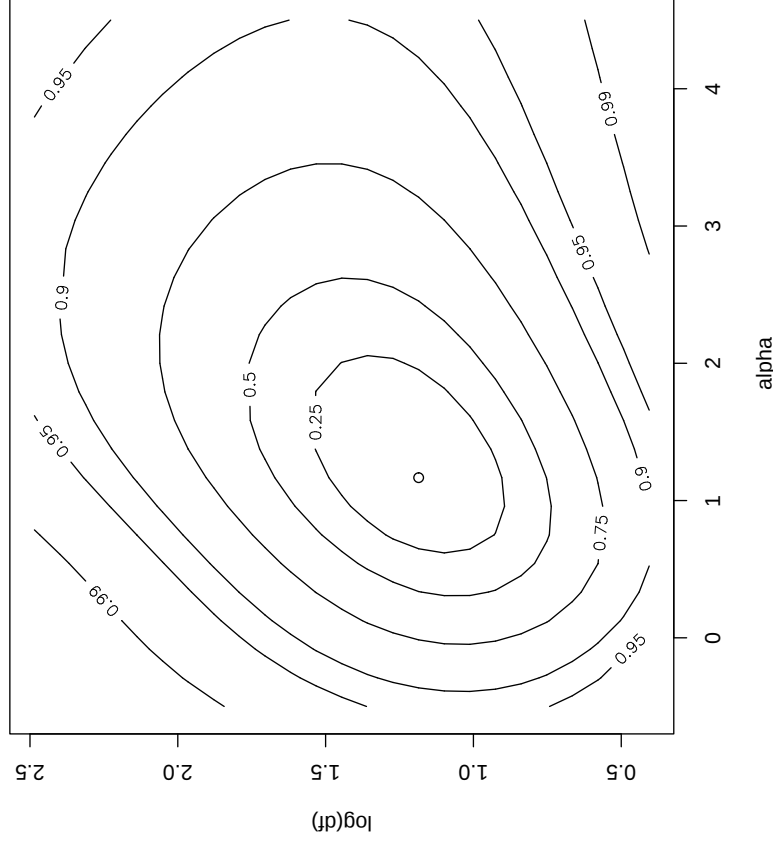
PP-plot for normal distribution



PP-plot for skew-t distribution



dataset: m.marietta
Profile deviance



PP-plot for Normal model and ST model

profile $-2 \log(\text{relative } L)$ for $(\alpha, \log \nu)$



Building expansions of arbitrary pdf's

SkSy class: $f(z) = 2 f_0(z) G\{w(z)\}$

- An arbitrary pdf $g(x)$ can be factorised as

$$g(x) = 2 f_\xi(x - \xi) G\{w_\xi(x - \xi)\}$$

where f_ξ is symmetric around 0, and w_ξ is odd.
The factorisation depends on the chosen ξ .

- Let $P_K(\cdot)$ denote odd polynomial of degree K . The class

$$2 f_0(x - \xi) G\{P_K(x - \xi)\}$$

is dense in SkSy-class (under some weak conditions).



Some applications



Use as a mathematical tool

not really an ‘application’, still useful

- to approximate a distribution (see above)
- local departures from normality, local efficiency, ...

[Salvan (1986), Durio & Nikitin (2003), Muliere & Nikitin (2004)]

- to produce ‘alternatives’ in simulation studies

[DiCiccio et al. (1997), Chu et al. (2001), etc.]

- ‘skew link’ for dicotomous response analysis

[Chen, Dey & Qi (1999)]



Connected areas

- Selected (non-random) sampling and Heckman model

response eqn: $y = \beta^\top x + \sigma \varepsilon_1, \quad \varepsilon_1 \sim N(0, 1)$

selection eqn: $z = \gamma^\top x + \varepsilon_2 > 0, \quad \varepsilon_2 \sim N(0, 1)$

$\implies f(y|x, z > 0) = \text{type ESN}, \delta = \text{corr}\{\varepsilon_1, \varepsilon_2\}$

[Copas & Li (1997)]

- stochastic frontier model

$$\begin{aligned} y_i &= f(x_i, \beta) - \sigma_1 |\eta_i| + \sigma_2 \varepsilon_i \\ &= f(x_i, \beta) + \omega \left(-\delta |\eta_i| + \sqrt{1 - \delta^2} \varepsilon_i \right) \end{aligned}$$

[Aigner, Lowell & Schmidt (1977), ..., Tancredi (2003+)]



Compositional data

$$(x_1, \dots, x_{d+1}), \quad x \in S_{d+1} \quad \left(\text{i.e. } x_j \in \mathbb{R}^+, \sum_{j=1}^{d+1} x_j = 1 \right)$$

- frequent sort of data in geology, among others areas
- popular approach: Aitchison ALR transformation to $\mathbb{R}^d$,
$$y_j = \log(x_j/x_{d+1}), \quad (j = 1, \dots, d),$$
enjoies a set of convenient formal properties
- then apply methods for multivariate Normal data to y_j 's
- use of SN_d in place of N_d preserves closure of suitable transformations of x on S_{d+1} : $a \oplus (b \otimes x)$ and subcompositions

[Mateu-Figueras (Ph.D thesis, 2003), Aitchison, Mateu-Figueras & Ng (2003)]



Applications in finance

- multivariate SN fits well into portfolio selection theory
use of skew- t allows fat tails

[Adcock & Shutes (1999), Adcock & Mead (2003), Harvey et al. (to appear)]

- incorporating skewness/kurtosis in components of
ARCH, GARCH and SV models

[Goria (thesis, 1999), Pietrobon (thesis, 2003), De Luca & Loperfido (2004), Cappuccio et al. (2004), Liseo & Loperfido (to appear)]



nearly there...



Some open problems

- a multitude of proposals (emphasis on probability side, less about statistics; what about their properties, connections, suitability for data analysis, etc.?)
- various issues in inference:
 - avoid some peculiarities of MLE
 - alternatives to MLE
- extend various methods from Normal to non-Normal case
- can approximate any distribution, but
 - what is 'best' type of approximation?
 - how to get approximation with prescribed properties?
- is all of this **really** useful?



Resources

References, some papers, software, etc. at

<http://azzalini.stat.unipd.it/SN/>

